

Academic Journal Pricing and Market Power: A Portfolio Approach

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Revised: November, 2000

Abstract

Library demand for academic journals is unique. Libraries maximize the usage of broadly-defined collections subject to a budget constraint, and the result is a demand for portfolios of titles. This paper describes the demand structure and develops a two-stage model of pricing. In this model publishers may find it profitable to increase the size of their journal portfolios via mergers. Using data for some one thousand biomedical journals, I estimate two empirical models including a structural model that is able to identify the impact of past mergers. The results indicate that the demand for journals is highly inelastic, that quality- and cost-adjusted price increases have been substantial over the past decade, and that past mergers have contributed to these price increases.

Acknowledgments

I would like to thank many of my former colleagues at the Department of Justice, including Craig Conrath, Renata Hesse, Aaron Hoag, Russ Pittman, David Reitman, Dan Rubinfeld, and Greg Werden, as well as Jonathan Baker, Cory Capps, George Deltas, Luke Froeb, Jeffrey Mackie-Mason, Roger Noll, Dan O'Brien, Richard Quandt, Lars-Hendrik Röller, Steve Salop and Margaret Slade; seminar participants at the Federal Trade Commission, Georgia Tech, SUNY Stony Brook, and Wissenschaft Zentrum Berlin; and participants at the meetings of the American Economic Association, the European Association of Research in Industrial Economics, the Southern Economics Association, and Western Economics Association. The Association of Research Libraries and its members, the National Library of Medicine, the Georgia Tech Library, and the Georgia Tech Foundation have provided invaluable assistance. Expert data support was provided by a large group of individuals, including Deena Bernstein, Claude Briggs, Pat Finn, Doug Heslep and Steve Stiglitz. Finally, I would like to thank the dozens of librarians and publishers who have provided me with important insights.

Despite their influence on the careers of economists, the production and pricing of scholarly journals have received scant attention from the profession.¹ By contrast, the issue of journal quality and scholarly research productivity have been studied in far greater detail (a search in the EconLit database using the term “journal” generates several dozen papers on this topic). Although there may be a number of reasons for this “imbalance” it is likely that the tenure process, combined with the low (if not zero) effective cost of journals on campuses have influenced our research agenda. In other words, while professors worry about their job security (publish *well*, or perish), others -- their librarians -- are charged with maintaining “free” access to all relevant journals. Of course, this pattern is observed not just in economics but across academic disciplines.

In recent years, however, easy access to *print* journals has been threatened.² Beset by persistent journal price inflation (especially in the so-called STM fields, or science, technology and medicine) and stagnant budgets, many university libraries have been forced to re-allocate dollars from monographs to journals, to postpone the purchase of new journal titles, and in some cases, to cancel titles. As a consequence, libraries often relied on inter-library loans to satisfy faculty demands. This situation and its possible causes has been studied at great length in the library science literature. With few exceptions, a consensus has evolved there which focuses on the growing importance of commercial publishers in the market for scholarly journals: *Over the past decade or more, commercial firms have aggressively raised prices at a*

¹ Two exceptions are Ordover and Willig (1978) and Chressanthis and Chressanthis, (1994).

² Increasingly, journals are available in both print and electronic versions, and for some new titles only an electronic format is available. The advent of electronic journals is *very* recent, however, and is unlikely to have influenced behavior during the sample period analyzed in this paper. See Tenopir and King (2000), chapter 15, for a discussion of these changes.

*rate disproportionate to any increase in costs or quality. This appears to be especially true for the largest commercial firms.*³ Although the analysis underlying these conclusions is generally not of the multivariate sort, it is suggestive enough to warrant further investigation.

This paper is the first to assess the merits of this “consensus.” Indeed, have changes in journal costs and quality accounted for most of the price inflation or has the exercise of market power by publishers played an important role? In addressing this question, the paper offers both theoretical and empirical support for the latter alternative. A model of journal pricing is proposed that reflects the underlying demand behavior of libraries. Although individual users are interested in just a handful of STM journals, libraries maximize the *usage* of broadly-defined collections, e.g. all biomedical journals, subject to a budget constraint. The result is demand for a *portfolio* of titles. In practice this means that libraries rank titles according to cost/use from lowest to highest and then select the largest set of low-ranked titles that they can afford. In other words, unlike most markets involving differentiated products, it is not appropriate to model demand as a discrete choice process. Rather, the typical library attempts to provide access to as many STM journals as possible through a combination of subscriptions and inter-library exchanges.

Given this portfolio demand, publisher pricing strategies are determined by the distribution of budgets and a title’s relative quality. Since all journals in a particular demand

³ See Tenopir and King (2000), Chapter 13, for a review of this literature. An alternative explanation for journal price inflation has been offered by Lieberman, Noll and Steinmuller in their 1992 working paper, “The Sources of Scientific Journal Price Increase,” Center for Economic Policy Research, Stanford University. They argue that entry by new titles over time has lowered circulation for existing journals, forcing the latter to raise prices to cover fixed costs. They estimate a supply and demand system for a set of journals and find that supply is downward sloping, consistent with this notion that individual titles exhibit scale economies. However, after controlling for this and other factors there remains a significant inflation residual that is unexplained by the model.

portfolio compete for the same budget dollars, relative quality determines demand for individual titles (if prices are equal, higher quality journals experience greater demand.). And the budget distribution influences whether, for example, high quality titles choose “low” prices and sell to most libraries or set “high” prices and sell only to the largest budget institutions. Furthermore, the pricing model predicts that in some cases firms controlling larger portfolios of journals have an incentive to charge higher prices, all else equal. Thus, past publishing mergers may account for some of the observed price increases.

To evaluate this and other conjectures, a unique dataset was assembled that includes cost, price and quality information for 990 biomedical titles as well as holdings information for these same journals at 194 biomedical libraries.⁴ This data is used to estimate a structural model that identifies the separate impacts of journal costs, quality, and publisher market power. The results indicate that the *firm-level* demand for journals is highly inelastic (for all firms, $|e_i| < 1$), that quality- and cost-adjusted price increases have been substantial over the past decade, and that past mergers have contributed to these price increases. The fact that journal demand at the firm level is inelastic is a sufficient condition for the exercise of market power. But the measured elasticities suggest that firms are not profit-maximizing. One possible explanation is that budgets for journals are sufficiently “soft” from year to year that price increases can be generally accommodated. Thus, each year publishers set new (higher) prices and libraries, trying to preserve their existing collections, respond by increasing budgets. Under such circumstances, the estimated, firm-specific demand elasticities *should* lie somewhere between zero and one, in absolute terms. This story can also account for the

⁴ This data collection effort began in 1998 while I was still employed by the U.S. Justice Department’s Antitrust Division. At that time, the Division was reviewing a number of proposed mergers between commercial publishers of STM journals, including (1) Reed-Elsevier, Wolters-Kluwer and Thomson; (2) Wolters-Kluwer and Waverly; and (3) Harcourt and Mosby.

estimated annual price increases. Finally, the merger-related price increases for the acquired firms' titles were substantial, about 25%. And yet the effects occurred in cases that would not raise antitrust concerns *even* if a portfolio approach to market definition was adopted, i.e. the combined portfolio market shares of the merging firms were less than 30%.⁵

These results raise a number of policy questions: 1. Since STM journal content is a public good (funded in most cases by tax dollars) with negligible distribution costs, does the performance of commercial STM publishing constitute a market failure? If so, do better alternatives exist? 2. Do antitrust authorities need a new paradigm for academic publishing and other portfolio-type markets? 3. How will the growing transition to electronic distribution affect the status quo? I briefly address these questions at the conclusion of the paper.

The paper is organized as follows. Journal demand and the implications for merger analysis are first discussed. Next, the journal pricing model is presented. Empirical evidence is then considered. Two empirical models and the underlying hypotheses are discussed, followed by a description of the data, and then the estimation results. Finally, I conclude by discussing the policy issues mentioned earlier.

Journal Demand and Merger Analysis

Most discussions of mergers start with market definition and thus demand. Because of our interest in related policy questions, consider how antitrust practitioners approach this problem. To assess the potential impact of a proposed merger, antitrust authorities define a

⁵ In recent years the DOJ has adopted a content-based approach to publishing mergers. Thus, unless substantial content overlap exists between merging firms no action is likely. Under this standard a merger of two firms controlling *all* titles in a particular market portfolio would be allowed so long as no content overlap was observed.

market just large enough so that a hypothetical monopolist in the market would find it profitable to raise prices some percentage amount above current prices. That is, market demand in the defined market is sufficiently inelastic to justify the price increase.

In scholarly publishing, experience as a user suggests that each *unique* journal title constitutes a distinct market for the purposes of antitrust analysis. For example, no one would argue that articles in *Brain Research* could be easily substituted for ones in the *New England Journal of Medicine*, much less those in the *American Economic Review*. If each title corresponds to an antitrust market, then owners of individual titles already have the capacity to achieve monopoly returns; a corollary is that mergers don't matter. Furthermore, even if markets are defined somewhat more broadly, say, to include titles whose content overlaps, the likelihood that two publishers (among dozens) would together control sufficient content (in a market with thousands of titles) to warrant antitrust scrutiny seems small.

These priors about market definition imply that demands for individual titles are unrelated. However, this intuition proved to be incorrect, at least for the major purchasers of STM journals, the libraries. Discussions with dozens of librarians revealed the following: purchase of academic journals by libraries is generally based on two factors — annual subscription price and expected usage. To assemble and maintain their collections, most libraries appear to construct a cost per use ratio for each title.⁶ Given a budget for a relevant academic field, e.g., biomedicine, they then proceed to rank journals from lowest to highest in that field according to this ratio, and identify a cutoff above which titles need to be canceled; conversely, if their holdings in the relevant field do not exhaust the budget, additional titles can

⁶ This claim is literally true for medical libraries; though other types of academic libraries may not be as precise in their processes, they appear to behave in similar fashion. In any case, this is an empirical question that is tested using the holdings data.

be purchased until the budget constraint is met. From year to year, as budgets and titles' usage change, collections are adjusted accordingly.⁷ Over the past decade or so the general trend is for increases in library budgets to lag journal price inflation; a consequence is that many libraries have been forced to re-allocate dollars from monographs to journals, to postpone the purchase of new journal titles, and in some cases, to cancel titles.

The most interesting aspect of library journal acquisition, of course, is that individual titles within a given field are considered simultaneously. So, for example, bio-medical libraries group titles from various sub-fields, e.g. neurology, biochemistry, clinical medicine, etc., into a single "portfolio" and broadly apply the cost per use criterion. Thus, titles compete with each other for budget dollars across an entire field, rather than across a narrow sub-field, as intuition might otherwise suggest (an intuition based on the perspective of the typical user of journal materials). Furthermore, since journal content is highly differentiated even within sub-fields, libraries try to provide access to as many STM journals as possible through a combination of subscriptions and inter-library exchanges.

Publishers' Pricing Strategies

Given this demand structure, how do publishing firms price their journals? First, there exist at least two types of journal companies— commercial and non-profit— and these have different strategies. In general, the latter are mostly intent on disseminating knowledge, whereas the former are primarily interested in profits. Here I assume that the non-profit firms

⁷ This type of constrained optimization problem is commonly referred to in the operations research literature as a "knapsack" problem. Note that in practice a library's budget constraint may be "soft"; it may be expanded slightly to accommodate the marginal journal that doesn't quite "fit" the budget. For an introduction to this issue in the OR literature see Daellenbach and George (1978).

set prices to cover costs (and are thus ignored in the analysis that follows).

A Model of Journal Pricing

Commercial journal publishers, like firms in any industry, will take into account the structure of demand and the likely strategies of competitors when setting prices. As described earlier, libraries – which constitute the bulk of demand for STM journals – attempt to purchase the most “usage” given their budgets for serials.⁸ In practice this amounts to ranking journals on a cost/use basis and identifying a threshold cost/use value above which titles are no longer added (the budget is exhausted).

More formally, suppose there exists T titles of varying usage or quality, u_t .⁹ Then given the price for each title, p_t , libraries can rank their journals on a cost/use basis,

$$p_{T-n}/u_{T-n} < p_{T-m}/u_{T-m} < \dots \quad \text{where } 0 < m, n < T$$

And given some budget B , a library can choose titles for its collection according to this ranking until B is fully allocated. The cost/use ratio for the last title added, p^*/u^* can be used to determine subsequent additions and cancellations of titles.

How are prices set in this demand environment? I consider two cases: one in which the number of libraries exceeds the number of possible budget levels, and a second case where each library budget is unique.

⁸ To simplify the analysis, the possibility of inter-library exchanges is ruled out. This assumption implies that, on a cost/use basis, the cost of inter-library loans exceeds the costs of all subscriptions. Also, the model assumes that libraries make no duplicate purchases, an assumption that is generally correct except for the highest-use journals.

⁹ To allow for the fact that journal usage will differ from library to library I permit the u_t to vary across institutions. However, I assume that the ratio u_t/u_s is constant in all libraries for any pair of titles, t and s . This assumption guarantees that cost/use rankings are identical across libraries.

Case 1: Budget Classes

Suppose that there are two types of library budgets, B_S and B_L , respectively, and $B_S < B_L$ (the analysis can be easily generalized to include additional budget classes). The corresponding population of each type is N_S and N_L (where $N_S + N_L > 2$). I assume that each of the T titles are sold by a separate publisher. No price discrimination is allowed, i.e. annual subscriptions are sold for a unique price. Common journal production costs consist of a fixed component, $F > 0$, and a marginal cost, $c = 0$.

We consider a two-stage game.¹⁰ In the first period, each of the T firms consider whether to target (through choice of content, advertising, etc.) all libraries or just those with large budgets. Once these sunk investments have been made, each firm takes into account the pricing strategies of firms that have made a similar marketing choice. To better understand this dynamic game, consider first the end-period pricing strategies.

Suppose z ($T - z > 0$) firms have chosen to target all libraries. For sales to be made to the entire set of buyers by each of these firms it must be true that

$$p_1 + p_2 + p_3 + \dots + p_z \leq B_S \quad (1)$$

Given this budget constraint, how do each of the z firms set its price? Given libraries' cost/use selection criterion, a "high" price raises a firm's profits so long as it's associated cost/use ratio lies below the threshold of each library. Consider the case where firms have set prices so that their cost/use ratios are all equal, that is,

¹⁰ Although equilibria exist in a one-stage version of this game, the set of parameters for which they exist is smaller. In any case, the two stage version is more realistic.

$$\frac{P_1}{u_1} = \frac{P_2}{u_2} = \dots = \frac{P_z}{u_z} \quad (2)$$

If (1) is also satisfied (with equality), then an increase in price by one firm results in the loss of its sales to at least the small budget libraries. A drop in price, on the other hand, has no effect on the firm's demand but lowers profits. Thus, when both (1) and (2) are satisfied, individual journal profits are maximized given the requirement that sales be made to all libraries. The result for the $T - z$ firms that target the large budget libraries is the same (note though, that in (1), B_S is replaced by $(B_L - B_S)$).¹¹

Are each of these budget-class constrained equilibria unique? Suppose (1) holds but not (2), i.e. at least two different cost/use ratios are observed. Then the lower ratio firms can increase profits with certainty by raising their prices a small amount. Libraries will respond by not purchasing from (some of) the higher ratio firms. Thus, in the second stage of the game, once firms have chosen a budget class, prices that satisfy (1) and (2) constitute a unique equilibrium. Using these two relationships, prices for firms targeting all libraries, and those targeting just large libraries, can be expressed, respectively, as

$$P_i = \frac{B_S * u_i}{\sum_{i=1}^z u_i}, \quad P_j = \frac{(B_L - B_S) * u_j}{\sum_{j=z+1}^T u_j}$$

The corresponding (gross) profits are

¹¹ Although cost/use ratios are identical for firms targeting the same budget class, this is unlikely *across* budget classes. In fact, such an occurrence is most improbable. As budget classes are added to the model, the number of unique cost/use ratios generally increases as well. This result corresponds to the empirical fact that cost/use ratios vary greatly.

$$\Pi_i = \frac{(N_S + N_L) * B_S * u_i}{\sum_{i=1}^z u_i}$$

and

$$\Pi_j = \frac{N_L * (B_L - B_S) * u_j}{\sum_{j=z+1}^T u_j}$$

As intuition might suggest, in a given budget class, price and profits are increasing in a journal's quality. On the other hand, the price and profitability of a journal is decreasing in aggregate class quality.

Given this end-period result, how do firms select their target budget class?

Since higher quality firms enjoy relatively more market power, one might expect these firms to target the most lucrative budget class. This in turn will be determined by the relative magnitudes of $(N_S + N_L)B_S$ and $N_L(B_L - B_S)$. For the remainder of this section of the paper, I assume that the former is larger. This assumption corresponds to the observation that in most journal markets the number of "small" libraries far exceeds the population of "large" ones, and that $B_L - B_S$ is not "too" large.

With this assumption, a potential equilibrium candidate for the first stage game involves the z highest quality firms targeting all libraries, and the remaining $T - z$ firms focusing on the large budget customers (an "ordered" equilibrium). For this to constitute a Nash equilibrium, no individual firm can wish to defect from its budget class. In the case of one of the z high quality firms this implies,

$$\Pi_i > \frac{N_L * (B_L - B_S) * u_i}{\left[\sum_{j=k+1}^T u_j \right] + u_i} \quad (3)$$

For the lower quality firms, the corresponding inequality is

$$\Pi_j > \frac{(N_S + N_L) * B_S * u_j}{\left[\sum_{i=1}^z u_i \right] + u_j} \quad (4)$$

Combining (3) and (4) yields the following equilibrium condition,

$$\frac{\left[\sum_{j=z+1}^T u_j \right] + u_i}{\sum_{i=1}^z u_i} > A > \frac{\sum_{j=z+1}^T u_j}{\left[\sum_{i=1}^z u_i \right] + u_j} \quad (5)$$

where

$$A = \frac{N_L (B_L - B_S)}{(N_S + N_L) B_S} < 1$$

The LHS inequality in (5) is simply a re-arranged form of (3); the RHS inequality is based on (4). For a Nash equilibrium to exist in this stage, (5) must hold for all firms. Furthermore, if (5) holds for the z th and T th journals, it is easy to demonstrate that (5) is satisfied for all firms. To see this, note that for the z th journal the LHS is at its *minimum* value ($u_z < u_i$ for all $i < z$); for the T th journal, the RHS ratio is at its *maximum* value ($u_T < u_j$ for all $j < T$). If (5) holds for these

two firms, then it is satisfied in all instances.¹² Thus, searching for an “ordered” equilibrium involves first checking that the RHS inequality holds for the Tth journal and then identifying the marginal high quality journal for which the LHS inequality is satisfied.

How do mergers affect outcomes in this simple model? There are a number of potential scenarios: mergers within budget classes, those across budget classes, and some combination of these first two cases. Consider the case of a within-class merger involving two high-quality firms. What pricing strategy does the merged firm adopt? As we noted earlier, a journal's profitability is decreasing in aggregate class quality. This suggests that the merged firm might benefit from raising the price of one of its titles enough to cause the small budget libraries to drop it and replace it with a lower quality title. This “jumping” between budget classes lowers the aggregate quality of firms targeting all libraries, and thus enhance the profitability of the merged firm's remaining general circulation title. The profitability of the “dropped” title may go up or down (if the pre-merger equilibrium was unique, then profits decline), depending on the model's parameters.¹³ The sum of these two components will determine the post-merger pricing strategy. If the net effect is positive, then the merger is harmful: the average quality of library collections decreases.

¹² To rule out other possible equilibria, it is necessary to check whether (5) would hold if one or more of the high quality firms exchanged places with lower quality firms, i.e. they switched budget classes.

¹³ Of course, a journal jump *between* budget classes influences the prices charged by other firms. In simulations of the merger scenario described above, the non-jumping journals experience modest price changes compared to the jump journals. The merged firm's high-quality, jump journal exhibits large price *increases*; the non-merger, low-quality, jump journal shows relatively large price *decreases*. This pattern persists as one increases the number of titles and budget classes. However, if the journal populations of particular budget classes are unchanged after a merger then the prices for those titles remain unchanged. Since it is likely that any observed merger will involve titles in *different* budget classes, it is possible that the merging firms' titles will jump in both directions, i.e. higher quality titles will jump “up” by increasing prices while lower quality titles will jump “down” by lowering their prices.

This result can be easily expressed using the above notation. Refer to the two high quality titles controlled by the merged firm as journals “1” and “2.” Suppose the title now targeted at large budget libraries (the “dropped” title) is journal 2. Its profits are then

$$\Pi_2^{post-merger} = \frac{N_L * (B_L - B_S) * u_2}{T [\sum_{j=z+1} u_j] - u^* + u_2} < \Pi_2^{pre-merger}$$

where u^* is the quality of the journal that now replaces journal 2 in small budget collections. I assume that the pre-merger equilibrium was unique. This implies that 2's post-merger profits decline.

Journal 1's profits are

$$\Pi_1^{post-merger} = \frac{(N_S + N_L) * B_S * u_1}{T [\sum_{i=1} u_i] + u^* - u_2} > \Pi_1^{pre-merger}$$

Journal 1's profits increase since aggregate quality in its budget class declines.

This strategy will be adopted if $\Delta\Pi_1 + \Delta\Pi_2 > 0$.

Case 2: Unique Budgets

Allowing for unique library budgets changes the analysis in subtle but important ways. Suppose N unique budgets, B_i , are distributed according to some continuous density function, $f(B_i)$, with support (B_L, B_H) and cdf, $F(B_i)$, where $F(B_L)=0$ and $F(B_H)=1$. Assuming the same two-stage game as before, analysis of the second stage in the game is simplified since each of the T firms target a different set of libraries in the first period.

Suppose this is true, then in the second period firms choose prices consistent with their first period choices. For example, consider an ordered equilibrium in which during the first period the highest quality firm, Z_1 , targets the largest population of libraries, with support (B_1, B_H) where $B_1 \geq B_L$; the next highest quality firm, Z_2 , targets a subset of this population with support (B_2, B_H) where $B_2 > B_1$, and so on. Then, in the second period, Z_1 sets a price, P_1 , equal to B_1 , Z_2 sets a price, P_2 , equal to $B_2 - B_1$, etc. Because of their first period choices, firms have no incentive to undercut (in a cost/use sense) other firms selling to larger library populations since undercutting has no impact on quantity sold but reduces profits. And any attempt to raise price reduces expected sales.

So how do the T firms choose their target population in the first period? Consider an ordered equilibrium where $T=2$. Given the distribution of N budgets, Z_1 needs to target a population set so that Z_2 has no incentive to “undercut” and target a larger set. In terms of second period prices,

$$P_2 N (1 - F(P_1 + P_2)) \geq P_2^* N (1 - F(P_2^*))$$

$$\forall P_2^* \text{ s.t. } P_2^* \leq (u_2/u_1) P_1 \quad (1)$$

where u_1 and u_2 correspond to the quality of Z_1 and Z_2 , respectively, and $u_1 > u_2$.

Similarly, given Z_2 's targeted population, Z_1 can't find it profitable to target a smaller, higher budget population:

$$P_1 N (1 - F(P_1)) \geq P_1^* N (1 - F(P_2 + P_1^*))$$

$$\forall P_1^* \text{ s.t. } P_1^* \geq (u_1/u_2) P_2 \quad (2)$$

For $T=3$, the number of relevant constraints increases to 6 (3×2); for $T=4$ the number is 12 (4×3), etc. Depending on journal quality, the distribution of budgets, etc., some of these

constraints may be redundant.

In a number of simulations, ordered equilibria were observed that resemble the actual data, e.g. high quality journals exhibit relatively low cost/use ratios while low quality journals have high cost/use ratios.¹⁴ Mergers between various firms in these simulations were often profitable. The impact on prices was generally asymmetric – the price of lower quality journals increased while those for the higher quality titles decreased. The impact on other firms' prices showed no clear pattern. Note that unlike the previous case, merged firms did not find it profitable to set prices that would result in a re-ordering of journal cost use rankings, i.e. journal jumps were not profitable.

Testing the Portfolio Theory

The Institute for Scientific Information (ISI) tracks citations in peer-reviewed titles for over 8,000 STM journals in various fields. Some guesses place the world population of academic journals close to 30,000. Not surprisingly, the number of publishers, both commercial and non-profit, is large as well. With respect to biomedical journals, ISI tracks titles published by at least 70 companies. Over the past decade a flurry of merger activity has been observed in the STM publishing market, particularly in the past two years. Since the latter half of 1997 alone, at least six major commercial publishers have been purchased by competitors. In addition, numerous small-scale transactions involving one or two journal titles occur every year.

Although these recent natural experiments will provide a rich empirical opportunity in the near future (once several years of post-merger prices are generated), two mergers that

¹⁴ In these simulations, each publisher initially owns a single journal. Library budgets are assumed to be distributed according to an exponential density function.

occurred earlier in the 1990s should shed some light on the likely impact of this ongoing merger wave. In 1991, Reed-Elsevier purchased Pergamon and its large portfolio of STM titles, including some 57 ISI-ranked biomedical journals. At the time, Elsevier's biomedical portfolio numbered 190 titles. During the same period, Wolters-Kluwer added Lippincott's 15 ISI-ranked biomedical titles to its collection of 75 biomed journals. Since that time both companies's portfolios have grown further. In 1998, according to ISI data, Elsevier's portfolio stood at 262 titles; Kluwer controlled 112 journals.

A. Empirical Models

Previous empirical studies of journal pricing have not attempted to assess the extent of market power in the academic publishing market. Cressanthis and Cressanthis (1994) specified a reduced form hedonic model to study the determinants of economic journal pricing. Their results suggest that prices are related to journal characteristics (e.g., prices are increasing in journal quality and size). Lieberman, et. al. (1992) estimated a supply and demand system using data for 225 ISI-ranked science journals. They find that supply is downward sloping, consistent with the notion that publishing is characterized by scale economies at the individual title level. Based on this evidence they argue (indirectly) that entry by new titles has lowered circulation for existing journals, forcing the latter to raise prices to cover fixed costs. Furthermore, they identify a significant inflation residual that is unexplained by the model.

Two empirical models are estimated here. First, to test whether libraries' acquisition strategies reflect a ranking of journals according to cost/use values, we estimate an exponential cumulative distribution function. In other words, are cost/use values and journal demand inversely related? Confirmation of this hypothesis provides support for our portfolio

approach to demand.

Second, we estimate a structural two-equation model that measures firm-specific demand elasticities and explicitly accounts for the possibility of increased market power due to past mergers. Recall that inelastic demand is a necessary condition for the exercise of market power by publishers. Evidence of merger-related price increases is consistent with a portfolio market definition as well as the type of strategic behavior implied by the pricing models. The relative accuracy of the two pricing models is determined by examining the cost/use rankings of journals pre- and post-merger. Journal jumping is predicted by the budget class pricing model; if publishers perceive budgets as unique then no significant re-ordering of journal rankings should be observed.

1. Exponential CDF

An examination of the Medical Library Association's membership list reveals that most member libraries maintain small to medium collections, reflecting their modest budgets. Major research collections constitute a relatively small fraction of this population. Thus, it is anticipated that demand for low cost/use journals should greatly exceed demand for high cost/use titles. The exponential density function and its corresponding cdf is appropriate in this context for testing whether higher cost/use values (observed in larger budget collections) and journal demand are inversely related.

After ranking journals according to their cost/use values, library budgets can be constructed from the observed price data. In other words, if library i in the sample subscribes to m titles from a population of n titles ($n > m$), these journals should be the m lowest cost/use journals, with i 's budget equal to

$$B_{im} = \sum_{i=1}^m p_i \cdot$$

The expected value for journal m 's demand can then be expressed as a product of the observed library population, N , and the cumulative probability of library budgets larger than or equal to B_{im} :

$$N \int_{B_{im}}^{\infty} \frac{e^{-B/a}}{a} dB = N e^{-B_{im}/a}$$

Taking logs, the estimating demand equation is

$$\log(q_m) = \log(N) - (1/\alpha) \cdot B_m + \epsilon$$

where ϵ is an iid disturbance term. To facilitate comparison across different years in the sample, e.g. 1991 and 1995, journals that had missing observations in either year were omitted. Furthermore, since demand for the relatively small number of general subject journals is distinct from the majority of more specialized titles, the former are excluded as well (see below for an explanation on how these general titles can be identified).

2. Structural Market Power Model.

The model of journal pricing presented earlier implies that publishers' prices are likely to exceed marginal costs. The extent of this market power is influenced by journal quality, library budgets, as well as the number and quality of competitive titles. Furthermore, under certain conditions it is profitable for a firm to raise its prices after an increase in portfolio size

resulting from a merger. The challenge in establishing the empirical importance of this latter possibility lies in isolating the influence of portfolio size on prices. Cost increases, quality improvements, changes in demand elasticity, etc. are all potential confounding factors. I specify a model below that is designed to identify the effects of mergers during the sample period. The empirical model borrows from the existing literature, especially Baker and Bresnahan (1988), Baker (1989), Rubinovitz (1993) and Parker and Röller (1997). None of these papers, however, consider the impact of mergers on pricing.

Consider the following inverse demand function:

$$(1) \quad P_i = f(Q_i, \mathbf{Q}, \mathbf{Y}_i) + \epsilon$$

where P_i is the price of journal i , Q_i is the number of annual subscriptions to journal i , \mathbf{Q} is the vector of subscriptions to other journals, \mathbf{Y}_i is a vector of exogenous factors that can shift demand, and ϵ represents random fluctuations in demand. Because of the large number of journals in the sample, it is necessary to reduce the dimensionality of (1). Furthermore, the main source of this problem, the vector \mathbf{Q} , is endogenous. To address both of these econometric issues, a single-dimensional, exogenous variable, Δu_i , is substituted. This variable is defined as $(u_i - u^*)$ where u^* is the mean value of journal quality in the sample. Δu_i is assumed to be exogenous, and acts as an instrument for \mathbf{Q} since sales of other journals are a function of their (relative) quality. After this substitution, the inverse demand can be written as

$$(1') \quad P_i = f(Q_i, \Delta u_i, \mathbf{Y}_i) + \epsilon$$

Using (1') a marginal revenue curve can be derived by differentiating revenue with respect to Q_i :

$$(2) \quad MR_i = P_i + Q_i f_Q + \epsilon$$

where f_Q (<0) is the partial derivative of the demand function with respect to Q_i . By equating (2) with the short run marginal cost of producing journal i , the following “quasi-supply” relationship can be defined:

$$(3) \quad P_i = -Q_i f_Q + c(\mathbf{Z}_i, \mathbf{W}_i) + \mu$$

where \mathbf{Z}_i and \mathbf{W}_i represent vectors of endogenous and exogenous factors, respectively, that may shift costs, and μ is a random cost shock. Typically, to allow for the possibility that market power is not fully exercised an additional parameter, θ , is introduced, so that (3) can be rewritten as:

$$(4) \quad P_i = -\theta_i Q_i f_Q + c(\mathbf{Z}_i, \mathbf{W}_i) + \mu$$

Nash behavior implies that $\theta_i = 1$; $\theta_i = 0$ corresponds to perfect competition. For a firm owning more than one title, it is expected that $\theta_i > 1$ for each title in the firm’s portfolio. Since journals are strategic complements, firms controlling multiple titles have an incentive to set quantities lower and prices higher than suggested by Nash behavior.

Note that the marginal cost function is independent of Q_i , an assumption that is generally accepted as true in journal publishing. This condition allows us to empirically identify an increase in market power due to mergers among publishers. When marginal costs are independent of Q_i (and efforts are made to account for other factors that may shift costs), there remain two other possible explanations for an observed increase in journal prices following a merger. One possibility is that the absolute value of journal i ’s demand elasticity has declined,

holding market power, or θ_i , constant. An alternative explanation is that θ_i has increased.

To simplify the estimation, the demand function for journal i is assumed to take the following form:

$$(5) \quad P_i = e^{a_0} e^{a_1 \Delta u_i} Q_i^{a_2} Y_i^{a_3} \dots Y_i^{a_N} e^e$$

Given this functional form, the derivative of demand with respect to Q_i equals

$$(6) \quad f_{Q_i} = a_1 P_i / Q_i$$

Substituting equation (6) into (4), equation (4) can be written as

$$(7) \quad P_i = -\theta_i \alpha_1 P_i + c(\mathbf{Z}_i, \mathbf{W}_i) + \mu$$

or

$$(8) \quad P_i = (c(\mathbf{Z}_i, \mathbf{W}_i) + \mu) / (1 + \theta_i \alpha_1).$$

If $c(\mathbf{Z}_i, \mathbf{W}_i) = e^{\beta_0} Z_i^{\beta_1} W_i^{\beta_2} e^\mu$ (where the Z_i and W_i are taken to be scalars), then after taking the natural log of both sides of (8) this equation becomes,

$$(9) \quad \ln P_i = \beta_0 - \ln(1 + \theta_i \alpha_1) + \beta_1 \ln Z_i + \beta_2 \ln W_i + \mu$$

Since $\theta_i \alpha_1$ is likely to be relatively small for our application, $\ln(1 + \theta_i \alpha_1) \approx \theta_i \alpha_1$,

and so (9) can be simplified to

$$(10) \quad \ln P_i = \beta_0 - \theta \alpha_1 + \beta_1 \ln Z_i + \beta_2 \ln W_i + \mu$$

This general form of the quasi-supply function is estimated below. However, θ_i , the primary parameter of interest cannot be identified. To overcome this difficulty, consider the following strategy. If $(\theta_i \alpha_1)$ changes after a merger, and (10) can be estimated using data from before and after a merger, then it is possible to identify the *change* in this product by estimating

$$(10a) \quad \ln P_i = \beta_0 - \theta_i \alpha_1 - (\theta_i \alpha_1)^* \cdot \text{Merger dummy} + \beta_1 \ln Z_i + \beta_2 \ln W_i + \mu$$

where *Merger dummy* is a dummy variable that takes on a value of one for post-merger observations associated with journal *i* and zero otherwise; $(\theta_i \alpha_1)^*$ is the post-merger change in this product. Therefore, estimating (10a) provides an estimate of price changes due to a merger, after controlling for changes in journal costs and quality. To determine whether the change in this product is due to a change in θ_i and/or α_1 it is necessary to estimate a modified version of (5) that provides values for α_1 and the change in this parameter, α_1^* .

The two equations to be estimated, the quasi-supply function and the demand function, have the following form:

$$(10a') \quad \ln \text{Price}_{ijt} = \beta_{ij} - (\theta_j \alpha_{1j})^* \cdot \text{Merger dummy}_{jt} + \beta_1 \ln \text{Papers}_{ijt} + \beta_2 \text{Time trend}_t + \mu_{it}$$

$$(5') \quad \ln \text{Subscriptions}_{ijt} = A + \delta_1 \cdot \Delta u_{it} + \delta_{1j} \ln \text{Price}_{ijt} + \delta_{2j} \ln \text{Price}_{ijt} \cdot \text{Merger dummy}_{jt} \\ + \delta_3 \cdot \ln \text{Citations}_{it} + \delta_4 \cdot \text{BH dummy}_i + \epsilon_{it}$$

The three subscripts, *i*, *j*, and *t*, refer to journal *i*, publisher *j*, and year *t*. Note that

$$\delta_{1j} = 1/\alpha_{1j}.$$

β_{ij} : firm *j*'s fixed effect, common to all journals *i* owned by *j*.

Subscriptions_{it}: the number of medical libraries that purchased annual subscriptions to journal

i in year t .

Price_{ijt} : the institutional annual subscription price for journal i in year t .

Citations_{ijt} : the number of citations received by articles published in journal i in years $t-4$ through t in year t .

BH dummy: a dummy variable that equals one if the journal is listed in the Brandon-Hill journal list, and zero otherwise.¹⁵ This variable indicates whether a journal is a general (=1) or specialized title (=0), and is thus an exogenous measure of circulation “potential.”

Papers_{it}: measured by the number of papers published by journal i in year t .

Time trend : a continuous variable that equals one if an observation is associated with 1988, 2 if the associated year is 1989, etc.

u_{it} : ($\text{Citations}_{it} - \text{MCITES}_t$) where the latter term in the parentheses is the mean value of citations in year t .

Demand for a journal is expected to be decreasing in price, increasing in citations and Δu , and increasing when journals are classified as a BH title. A publishers' journal price is expected to be decreasing in the subscription count, increasing in the number of papers published, and increasing over time.

Three variables in the model are treated as endogenous - *Price*, *Subscriptions*, and *Papers*. Publishers may change a journal's paper count for any number of reasons – increasing (or decreasing) interest in a subject, as part of a strategy to enhance quality, etc. – and can do so in a timeframe of a year or less. Therefore, it is reasonable to treat *Papers* as an endogenous factor. To improve the precision of the 2SLS estimates, two excluded exogenous variables are used as instruments: $\log \text{Age}$ and $(\log \text{Age})^2$, where $\text{Age} = [\text{Year of}$

¹⁵ See Brandon and Hill (1997).

observation - journal i 's year of first publication], e.g. [1995 - 1945]. A journal's age is a good instrument for *Papers* in particular because older journals tend to publish more articles.

The effects of two mergers are evaluated. The firms involved were Reed/Elsevier and Pergamon, and Wolters/Kluwer and Lippincott. Both combinations occurred during 1990-91.¹⁶ For these 4 publishers, $Merger\ dummy_{jt} = 1$ for the period 1992-98. Furthermore, in (10a'), separate intercepts are estimated for each publisher; for each firm involved in a merger a separate $Merger\ dummy_{jt}$ coefficient is estimated. Similarly, in (5'), separate $Price_{ijt}$ coefficients (corresponding to the demand elasticities) are estimated for each publisher; and again, for each firm involved in a merger a separate value for α_{ij}^* is estimated.

B. Data

For the period 1988-98, the DOJ collected publisher and price data for some 3000 journals, and holdings information from various libraries. This data has been supplemented with additional information extracted from the Institute for Scientific Information's (ISI) Journal Performance Indicators database (JPIOD). This database allows us to calculate annual citation rates for individual journals¹⁷; JPIOD also includes the number of papers published annually by each journal during the sample period.

This paper's empirical discussion is focused on a subset of these several thousand journals, namely, biomedical titles. The reasons for this choice are several. First, based on our discussions with various librarians, biomedical libraries are most likely to evaluate their purchases using the portfolio approach described earlier; furthermore, these libraries typically

¹⁶ To best evaluate the impact of these mergers, only journals that existed *prior* to 1989 are included in the estimation. This is the vast majority of observations in the sample.

¹⁷ Journal citations are used as a proxy for actual usage in libraries.

make no distinctions among various biomedical disciplines, permitting us to consider all biomedical titles as part of a single, large portfolio. Finally, practical considerations, including the fact that biomedical holdings data is reported in a relatively standard fashion, encouraged our initial focus on this subset of titles.¹⁸

During the sample period, almost two thousand ISI-ranked biomedical journals were published; price data were available for about 1800 of these titles (complete time series were available for most but not all journals). Of this latter group, almost 1400 were published by organizations with at least three ISI-ranked titles. For the analysis presented here, only journals sold by commercial firms with portfolios consisting of ten or more titles were considered (thus excluding journals distributed by small private publishers as well as the non-profits). Complete holdings data for 194 U.S. medical libraries were collected, representing in aggregate some 60,000 subscriptions to ISI-ranked journals; the libraries were randomly selected from among Medical Library Association members (the MLA's membership numbers approximately 1500). Libraries of all sizes are represented in the sample, some holding less than ten subscriptions, while others report collections exceeding 1,300 titles.

The sample period, 1988-1998, is useful in at least two respects. First, it is sufficiently long to assess whether inflation continues to plague the journal market (and dovetails nicely with Lieberman, et. al.'s (1992) sample, which includes data from 1978-88). Second, as described above, the period contains a number of natural experiments, i.e., publishing mergers, that enables us to identify the impact of mergers on pricing. Growth via merger

¹⁸ Unlike most fields, biomedical scholars enjoy the use of the National Library of Medicine's central database that contains information on several thousand medical collections. Although this data source offered substantial benefits with respect to the initial phase of data collection, the data was not ideally organized for analysis purposes. One of the major difficulties was that much of the data - some 25% -- was too idiosyncratic for data processing; as a consequence several hundred additional hours of manual effort were required to transform the data into usable form.

should be distinguished from organic internal growth arising from the introduction of new titles. The latter may produce benefits (such as coverage of emerging fields of study) that helps to offset any intentional competitive harm; harm associated with acquisitions, on the other hand, is less likely to be balanced by substantial benefits – journals are simply reshuffled, and any fixed cost savings seem to be small.¹⁹

Descriptive Statistics

Using the ISI-defined biomedical portfolio and the corresponding library holdings we can calculate the actual size of various commercial publishers' journal portfolios as well as the observed sample portfolios. Table 1 reports this information using both sources. It is clear from this table that significant variation in portfolio size exists in the industry. Note that, based on the ISI numbers, the proposed 1998 merger between Reed/Elsevier, Wolters/Kluwer and Thomson would have affected about 42% of the biomedical titles owned by large commercial publishers.

Table 2 presents information on average price, citations, cost/use (price/citation), and number of papers published for each publisher in the years 1988 and 1998. Though prices, citations and paper counts generally increased during the period, the rate of change for prices was far more striking, resulting in higher cost/use numbers by the end of the period. For example, Elsevier's average journal price more than tripled during the period, while the corresponding citation and paper counts increased less than 25%.

Table 3 provides average circulation rates for titles by publisher in 1988 and

¹⁹ Furthermore, if publishing mergers do result in cost savings, economic theory implies that post-merger prices should *decline*, everything else equal.

1998.²⁰ Given that nominal prices for increased dramatically over the sample period, the apparent inelasticity of demand indicated by these numbers is remarkable. It suggests that library serials budgets increased sufficiently during the period to absorb most of the price increases. At the same time, these numbers provide indirect support for the model of journal pricing presented earlier in the paper: all else equal, if library budgets increase, firms have an incentive to proportionally raise prices.

Estimation Results

The estimation results for the exponential cdf, and structural market power are reported in Tables 4 and 5, respectively. In Table 4, the results for the exponential cdf model are consistent with expectations. The *budget9X* parameters are negative and significant in both 1991 and 1995. The results suggest that higher cost/use journals are purchased by few libraries. For example, in 1991, the marginal journal for a \$100,000 budget library has a cost/use value equal to about 0.22 and, using the parameter estimates, is held by about 30 of the 194 libraries in the sample. The marginal journal for a \$200,000 budget library has a cost/use value equal to about 0.59, and is held by some 17 libraries. The smaller parameter value in 1995 reflects the combined effect of low demand elasticity and substantial price inflation over the 4 year period. A library that purchased the complete set of titles in both years experienced about a 70% increase in outlays, from approximately \$292K to \$498K.²¹

Turning to the structural model results, Table 5 reports only the parameters of interest,

²⁰ These numbers exclude titles that commenced publication after 1988. Including these newer titles would tend to lower the reported 1988 figures relative to the later 1998 numbers.

²¹ Remember that this “complete set of titles” is a subset of the sample analyzed in the structural model.

e.g. the various firm-specific intercept terms in the quasi-supply function are omitted. Note also that in the demand equation, the price coefficients for Pergamon and Lippincott are estimated relative to the corresponding estimates for their merger partners.

Parameters for the *Time Trend*, *In Papers*, *Brandon Hill*, u , and *In Citations* variables are all precisely estimated and have the expected signs. The *Time Trend* estimate implies that, after controlling for changes in quality and costs, publishers increased annual journal prices some 140% over the 1988-98 period (over the same period the CPI increased by 37%). Demand for titles included in the Brandon-Hill list is on average 90% larger than demand for more specialized journals. The u parameter suggests that as a journal's quality improves relative to the average value in the sample, demand increases. All else equal, higher values for *In Papers* and *In Citations* increase equilibrium prices and shifts demand outward, respectively.

Demand is apparently very inelastic. The reported parameters indicate that the merging firms' demand elasticities were all greater than -0.36 (none of the omitted elasticity estimates were less than -0.50). Since profit-maximizing firms normally operate in the elastic region, how can we account for this apparent contradiction of economic theory? One possibility is that budgets for journals are sufficiently "soft" from year to year that price increases can be generally accommodated, i.e. demand is inelastic for observed prices. Thus, each year publishers set new (higher) prices and libraries, trying to preserve their existing collections, respond by increasing journals budgets. This story is consistent with numerous librarians' experiences and with what some publishers have privately acknowledged.²² Under such circumstances, the estimated, firm-specific demand elasticities *should* lie somewhere between

²² According to one former publishing executive, "If we didn't raise our prices each year, our competitors would grab the surplus dollars available from our customers."

zero and one, in absolute terms.²³ Note that this story can also account for the estimated annual price increases.

Did the two publishing mergers earlier in the decade enhance the participating firms' market power? With respect to the Elsevier-Pergamon transaction the answer seems clear. Since the *Price•Merger dummy* parameters in the demand equation for these two firms are equal to zero, we can interpret the corresponding *Merger dummy* coefficients in the quasi-supply equation as pure, merger-induced market power effects. Post-merger, Elsevier journal prices increased about 5.2% and the former Pergamon titles experienced a 27% increase. This asymmetry is observed in the Kluwer-Lippincott merger as well. Post-merger, the former Lippincott titles, experienced a 30% price increase while the Kluwer prices dropped slightly. However, in this case the Lippincott price increase is not solely a consequence of enhanced market power. The Lippincott *Price•Merger dummy* estimate implies that demand for Lippincott titles became slightly more inelastic in the post-merger period, contributing at least partially to the observed 30% price increase.²⁴

Are these results consistent with either of the two pricing models? In both cases, the

²³ In a single period game, each publisher would attempt to forecast the size of journal budgets, and set prices so that its average absolute demand elasticity was close to one. However, in a multi-period context, with budgets increasing each period, a firm's pricing strategy changes. It is possible to show that firms will set prices so that absolute elasticities in each period lie between zero and one. The intuition is that lowering the price (and thus the absolute elasticity) in each period preserves future sales, and combined with budget growth, raises total profits.

²⁴ The relative contributions depend on the prior beliefs held regarding the pre- or post-merger value for θ . In the case of Lippincott, the estimated parameter values imply that if the post-merger value for θ equals 0.21 or less then the merger increased market power; if not, then post-merger a *decline* in market power is offset by an absolute decline in demand elasticity. Since it is unlikely that a merger would diminish a firm's market power, one should probably conclude that θ is below 0.21. This is also consistent with the fact that the estimated demand elasticities are below one. For some (public relations?) reason, publishers are not setting prices high enough to achieve absolute demand elasticities greater than one.

type of asymmetric price changes observed here are possible. For example, since the average Elsevier title exhibits higher quality and a higher cost/use ranking than the representative Pergamon journal, the budget class model might predict significant price declines for Elsevier in addition to the estimated increases for Pergamon titles. However, in each transaction the quality and cost/use rankings of the affected titles varied greatly. And when substantial variance in journal characteristics is assumed for the merging firms, this same model can generate modest price changes for one firm and large changes for the other.

Nevertheless, there does exist one robust difference between the two pricing models. Recall that in the budget class model the potential for profitable journal jumping is the underlying rationale for mergers. As a result, the prices of some titles should change enough so that their respective cost/use rankings shift. On the other hand, if budgets are unique journal jumping should not be observed. Table 6 reports average cost/use rankings for 692 journals divided into five sets, based on pre-merger ownership status: Elsevier, Pergamon, Wolters-Kluwer, Lippincott and all other titles. Three sets of journal rankings are constructed based on different cost/use ratios: 1. 1990 prices/1990 citations, 2. 1990 prices/1996 citations, 3. 1996 prices/1996 citations. The first and third ratios reflect journal cost/use rankings for the pre-merger year 1990 and the post-merger year 1996, respectively. The second ratio captures the contribution of changes in citation levels to the post-merger rankings, holding prices at their 1990 levels.²⁵

Comparison of the first and third ratios across the five sets of titles reveals that changes in the merging firms' journal rankings occurred. In each case, these shifts correspond

²⁵ These rankings are not sensitive to the choice of pre- and post-merger dates. The number of journals employed in this comparison is less than those in the full sample because pre-merger prices were missing for some titles.

to the merger effects estimated in the structural model. Average rankings for the acquired firms, Pergamon and Lippincott, increased substantially. Elsevier's ranking increased slightly while Wolters-Kluwer's value declined considerably. The average ranking of all other titles was essentially unchanged between 1990 and 1996. However, if we attempt to control for the impact of (exogenous) changes in citation levels during the six year period we find that the merger-induced shifts in cost/use rankings are typically more modest in magnitude. For the acquired firms, about 60% of the observed shifts up in ranking appear to be merger-related. The decline in Wolters-Kluwer's ranking is due entirely to changes in citation levels. The exceptions here are Elsevier (and other titles); the observed small increase (decrease) in its ranking is underestimated by using 1990 citation levels as a reference point. Nonetheless, it seems that journal jumping is an important consequence of journal mergers. Publishers appear to behave as described by the budget class model.²⁶

Policy Implications and Future Directions

A. Market Failure?

Efficient pricing is not sustainable in the declining average cost environment of academic publishing. This begs the question, how does the performance of commercial publishers compare to a second-best break-even standard? Our analysis suggests that prices far exceed *marginal* costs, but do they exceed *average* costs? One way to assess this question is to examine the pricing of comparable non-profit titles; presumably non-profit publishers set prices closer to if not equal to average costs. If the latter prove to be cheaper,

²⁶ Of course, this model's assumption that classes of identical library budgets exists is unlikely to be satisfied in a literal sense. However, there may be a sufficient number of similar budgets to warrant the type of behavior implied by the model.

then scholars have a real alternative for disseminating scholarly information in a more efficient fashion. Though a comprehensive analysis of non-profit journals is beyond the scope of the present paper it is useful to report some initial qualitative results. When descriptive statistics are generated for both commercial and non-profit, ISI-ranked biomedical journals, the discrepancy in average prices and quality for the two groups is striking. For example, if we compare titles of similar vintage we find that the average non-profit subscription price is between fifty to seventy-five percent less than the commercial titles' rates. At the same time average citation rates for the non-profit journals greatly exceed those of the commercial publishers' in most instances, sometimes by a factor of five.²⁷ Among commercial journals, prices and citations are positively correlated. Thus, the substantially lower prices of comparable non-profit titles suggests that commercial publishers are setting prices well in excess of average costs. Despite this apparent superiority²⁸, the population of ranked non-profit titles is far smaller than that of the commercial journals, 148 versus 1032. Has the lucrative journals market induced too much entry or have the non-profits been too slow to exploit emerging research areas? Although this question deserves further attention it seems clear that the two distinct publishing models exist, each successful in their own way.

²⁷ These two comparisons rely on titles with a first publication prior to 1988. Titles are aggregated according to the decade of initial publication, going backward from 1987. For younger titles, the average prices for the two groups are similar but the non-profit citation rates are about five times larger.

²⁸ Some of this citation gap may be due to the more general subject matter of many non-profit titles, compared to the niche strategy of some commercial journals.

B. Antitrust Paradigms

When the proposed 1998 merger between Reed Elsevier and Wolters Kluwer collapsed, opposition from antitrust authorities in Europe and the U.S. was cited as a primary cause. Although no formal complaints were filed by agencies on either side of the Atlantic, regulators had sent a variety of signals indicating their serious concerns. Negotiations with the European Union had progressed the farthest and it appeared that the proposed deal would proceed only if the parties agreed to significant divestitures. It was widely reported at the time that the EU's preferred set of divestitures upset the financial logic of the merger and resulted in its demise.

What is interesting here is that the EU's main focus was *not* on academic journals, but rather legal publishing (in Europe), and that its theory of anti-competitive harm was based on traditional antitrust principles, i.e., excessive overlap in content (and therefore similar to the DOJ's approach to the 1996 merger of legal publishers Thomson and West). The U.S. focus, of course, was far different, in part because European legal publishing was not germane and because the model of harm relied upon was novel.

Though one can only speculate on how a U.S. antitrust case might have proceeded, it is clear that the combined Reed-Elsevier/Wolters-Kluwer entity would have controlled large journal portfolios in a number of broad fields, including biomedicine. Assuming that these broad fields constituted antitrust markets, some of these portfolios would have crossed the U.S. Government's concentration threshold (based on the Antitrust Guidelines) with shares in excess of 30-35%. Based on the results discussed here, such a merger would have resulted in substantial price increases over time. If the U.S. had filed a complaint and had been successful with this market definition, an important *legal* precedent would have been set, one that would have made it easier to employ a portfolio theory in mergers involving combined

market shares less than the threshold, e.g. the subsequent Wolters-merger of Wolters-Kluwer and Waverly merger, *and/or* a large firm buying a relatively small portfolio of journals. The recent reluctance of the Antitrust Division to oppose several mergers in the publishing industry can be partially attributed to “insufficient” market shares. However, since many future deals are likely to be relatively small in scope, opposition to journal mergers will need to adopt novel approaches in the definition of both markets and concentration thresholds.²⁹

C. A Digital Future

Scholarly journals render at least three services: research communication, archiving, and quality certification. Digital technology offers the potential to transform the first two by providing instantaneous access to current and past research. With modest investments in computer hardware and software, global scientific communities can dramatically lower the costs of exchanging information.³⁰ Though these innovations might seem to threaten the future of the traditional journal, the latter’s role as a quality filter may be sufficient to preserve its existence, albeit in modified form. Although it is possible to conceive of new mechanisms for evaluating journal quality, e.g. measuring the number of “hits” generated by a journal website, it seems likely that the existing expert-based system for assessing new research will survive.³¹

²⁹ To avoid future antitrust scrutiny the Elseviers of the journal publishing world are likely to grow by adding relatively small numbers of journals at frequent intervals. If pursued diligently, this stealth strategy can be just as successful as any blockbuster merger.

³⁰ The Los Alamos physics server is perhaps the best example to date of this digital future (go to <http://xxx.lanl.gov/>). This website, funded by US government sources, has become the standard method of exchange for physics working papers.

³¹ One important justification for this claim is that professional advancement within (academic) institutions relies on and supports the existing approach to quality assessment.

Commercial publishers have begun to exploit these new opportunities by bundling their individual journal titles and providing libraries with electronic access to “article databases.”³² In doing so, the economics of commercial publishing may change in (subtle) ways. Portfolio size will still matter, but the number of “journals” may matter less than the total article population. Digital technology will make it feasible to control, monitor and price access in new and myriad ways, suggesting that sophisticated price discrimination schemes could be observed someday. The prospect of bundling and price discrimination, of course, will inevitably raise antitrust issues. A few, large portfolios reduce transactions costs for libraries yet have the potential for influencing new entry as well as pricing.

³² For example, Elsevier’s database product, *ScienceDirect* (www.sciencedirect.com), contains articles from its more than 1100 peer-reviewed journals in all STM disciplines. To gain access to the entire database or some customized subset, a library is required to maintain its Elsevier paper subscriptions. The access price is typically calculated as a percentage markup on the library’s Elsevier “paper budget.”

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TABLE 1

ISI-Ranked Medical Titles from Major Commercial Publishers, 1998

	# of titles published	# of observed ISI titles	%
Blackwell	112	99	0.88
<u>Churchill Livingstone</u>	17	12	0.71
Elsevier	262	225	0.86
Harcourt	118	109	0.92
Karger	45	39	0.87
<u>Mosby</u>	27	25	0.93
<u>Plenum</u>	22	20	0.91
Springer	99	87	0.88
Taylor	19	16	0.84
<u>Thomson</u>	41	36	0.88
<u>Waverly</u>	37	35	0.95
Wiley	78	70	0.90
Wolters-Kluwer	112	98	0.88
Totals	989	871	0.88

Notes:

- observed data based on holdings for 194 medical libraries, during 1988-98 period
- major firms are those with at least 10 ISI-ranked biomedical journals
- underlined firms were acquired in mergers with other firms in the list during 1997-1998: Churchill and Mosby by Harcourt, Plenum, Thomson, and Waverly by Wolters-Kluwer

TABLE 2

Selected Descriptive Stats, Avg. Values by Publisher

	1988				1998			
	Price	Cites	Cost/Use	Papers	Price	Cites	Cost/Use	Papers
Blackwell	193	1575	0.40	123	508	2652	0.55	156
Churchill Livingstone	183	1726	0.26	103	721	2821	0.62	146
Elsevier	482	3477	0.36	179	1548	4222	0.78	204
Harcourt	209	3713	0.18	164	518	5294	0.34	171
Karger	321	893	0.59	86	711	935	1.01	79
Mosby	100	4071	0.07	248	241	5369	0.15	269
Plenum	233	1352	0.25	92	759	1733	1.86	121
Springer	481	2268	0.44	141	1057	2386	0.84	153
Taylor	259	759	0.48	74	658	572	1.67	55
Thomson	207	1210	0.46	92	733	2788	0.45	140
<i>Waverly</i>	119	3171	0.10	188	277	5770	0.16	237
Wiley	333	2205	0.38	128	1409	3338	1.10	145
Wolters-Kluwer	176	2535	0.19	154	504	3519	0.52	153
Unweighted Avgs	253	2227	0.32	136	742	3184	0.77	156

Notes:

-- numbers based on journals that commenced publication prior to 1989,
and had ≥ 100 cites in 1988 or 1998.

TABLE 3

Avg. Circulation for ISI-Ranked Journals by Publisher

	1988 (# subscribers)	1998 (# subscribers)
Blackwell	31.72	30.16
Churchill Livingstone	34.00	31.20
Elsevier	30.08	27.92
Harcourt	50.51	53.23
Karger	28.81	22.77
Mosby	94.50	96.55
Plenum	27.61	22.89
Springer	21.60	19.03
Taylor	11.67	12.08
Thomson	13.50	19.42
<i>Waverly</i>	61.67	63.41
Wiley	24.41	23.51
Wolters-Kluwer	41.62	42.28
Unweighted Avgs	36.28	35.73

Notes:

- numbers based on holdings for 194 medical libraries, during 1988-98 period
- All titles commenced publication prior to 1989

Table 4, Exponential CDF

1991 Data

Dependent Variable: Ln Subscriptions

Variable	Estimate	Error	Prob > T
Intercept	4.0227	0.0634	0.0001
Budget	-6.0520E-06	3.5000E-07	0.0001

of observations: 658

Adj R-squared: 0.3095

1995 Data

Dependent Variable: Ln Subscriptions

Variable	Estimate	Error	Prob > T
Intercept	3.8783	0.0614	0.0001
Budget	-3.5490E-06	2.1000E-07	0.0001

of observations: 658

Adj R-squared: 0.2922

Table 5, Structural Market Power Model

Quasi-Supply Estimation Results

Dependent Variable: Ln Price

Variable	Estimate	Error	Prob > T
Elsevier Merger dummy	0.0507	0.0382	0.1838
Pergamon Merger dummy	0.2384	0.0691	0.0006
Wolters Merger dummy	-0.0418	0.0573	0.4663
Lippincott Merger dummy	0.2612	0.1184	0.0274
Ln Papers	0.3686	0.0081	0.0001
Time Trend	0.0874	0.0027	0.0001

of observations: 7588

Adj R-squared: 0.5130

Demand Estimation Results

Dependent Variable: Ln Subscriptions

Variable	Estimate	Error	Prob > T
Ln Price - Merger dummy, Elsevier	-0.0034	0.0079	0.6693
Ln Price - Merger dummy, Pergamon	-0.0034	0.0139	0.8066
Ln Price - Merger dummy, Wolters	-0.0268	0.0132	0.0427
Ln Price - Merger dummy, Lippincott	0.0675	0.0294	0.0219
Ln Price, Elsevier	-0.3552	0.0384	0.0001
Ln Price, Pergamon, relative to Elsevier	0.0337	0.0252	0.0497
Ln Price, Wolters	-0.3212	0.0425	0.0001
Ln Price, Lippincott, relative to Wolters	0.0495	0.0252	0.0497
Brandon Hill dummy	0.6265	0.0496	0.0001
$\Delta \hat{U}$	1.9163E-05	2.1940E-06	0.0001
Ln Citations	0.2279	0.0084	0.0001

of observations: 7588

Adj R-squared: 0.4065

Table 6, Cost/Use Rankings

Journal Group	Avg. Ranking		
	1990/1990	1990/1996	1996/1996
Elsevier (119)	365	346	369
Pergamon (45)	341	366	403
Wolters-Kluwer (48)	363	316	319
Lippincott (14)	140	150	166
All Other Titles (466)	347	354	344

of observations contained in parentheses